GLM point process for spike trains

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Review of generalized linear point-process models

Neurostats club 2015

September 21, 2015

GLM point process for spike trains

This will be quick No formal derivation No proofs

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Truccolo et al. 2005 figures and equations

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Truccolo et al. 2005

- Derivation of point process GLM
- Intrinsic and ensemble history models
- Maximum likelihood estimation
- Non-GLM conditional intensity model
- KS test and time rescaling
- Residual analysis
- Model selection
- Decoding

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Truccolo et al. 2005 figures and equations

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- Take NEUR2110 with Wilson Truccolo in the spring!

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Truccolo et al. 2005

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• Take NEUR2110 with Wilson Truccolo in the spring! If there is time:

• Closed form approximation, double crossvalidation, regularization, and regularization paths

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Linear model / multiple linear regression

$$y = \beta_o + \beta_1 \cdot x_1 + \beta_2 \cdot x_2 + \dots = \beta_o + \sum_{i=1}^N x_i \beta_i = XB + \beta_o$$

- y: response/dependent variable being modeled (when multivariate, often a column vector by convention)
- X: "design matrix" matrix of observations. Conventionally, each row is a realization and each column is a feature/variable
- B: Coefficient or parameter vector
- β_o : Constant term.

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Nonlinear features OK

Features (x₁, x₂, ...) can be any function of recorded data.
 E.g. the following model is common for phase/direction tuning (used in eqn. 10 in Truccolo et al. 2005)

$$y = \beta_o + A \cdot \cos(\varphi - \varphi_o)$$

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• A: amplitude parameter

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$$y = \beta_o + A \cdot \cos(\varphi - \varphi_o)$$

- A: amplitude parameter
- φ : the observed phase or direction

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- φ_o : preferred phase parameter

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$$y = \beta_o + A \cdot \cos(\varphi - \varphi_o)$$

- A: amplitude parameter
- φ : the observed phase or direction
- φ_o : preferred phase parameter
- Can be written in a form that is linear in parameters

$$y = \beta_o + A\cos(\varphi_o)\cos(\varphi) + A\sin(\varphi_o)\sin(\varphi)$$

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$$y = \beta_o + \beta_1 \cos(\varphi) + \beta_2 \sin(\varphi)$$

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- Truccolo et al. 2005:
 - "neural spike trains form a sequence of discrete events or point process time series"

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- Truccolo et al. 2005:
 - "neural spike trains form a sequence of discrete events or point process time series"
 - "standard linear or nonlinear regression methods are designed for analysis of continuous-valued data and not point process observations"

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Truccolo et al. 2005 figures and equations

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 - "standard linear or nonlinear regression methods are designed for analysis of continuous-valued data and not point process observations"
- Smoothing or binning can alter the structure
- GLM Point-process models directly model spike trains without these drawbacks

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- Consider a homogeneous Poisson process with rate λ .
 - expected # observations in time window Δ is $\lambda\Delta$

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- Consider a homogeneous Poisson process with rate λ .
 - expected # observations in time window Δ is $\lambda\Delta$
- For inhomogeneous Poisson process rate varies time $\lambda(t)$
 - # observations from t to $t + \Delta$ is $\int_t^{t+\Delta} \lambda(t) dt$

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- In a point process model we predict conditional intensity based on measured covariates $\lambda(t|X(t))$

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- $\lambda(t)$ is called the "intensity function"
- In a point process model we predict conditional intensity based on measured covariates $\lambda(t|X(t))$
- The Poisson GLM point process framework models conditional intensity functions of the form

$$\lambda(t|X(t)) = \exp\left\{\mu + XB\right\}$$

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Fitting GLM point-process models

- In practice spiking time series are discretized at some resolution Δ

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 - With a discrete approximation of the conditional intensity $\lambda_t = (\lambda_1, \lambda_2, \ldots)$

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mathworks.com/help/stats/glmfit.htmlt

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• Python scikits.statsmodels.GLM

statsmodels.sourceforge.net/stable/glm.html

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• I typically use gradient descent

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GLM point process model for single unit spiking with ensemble history

 $\log[\lambda(t|X(t))\Delta]=\mu$

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GLM point process model for single unit spiking with ensemble history

 $log[\lambda(t|X(t))\Delta] = \mu$ + intrinsic history

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GLM point process model for single unit spiking with ensemble history

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Intrinsic history filter

• Theoretically, model the whole history

$$\lambda_t(t|\theta, y_{t-1}, y_{t-2}, ..., y_1) \propto \exp\left\{\sum_{\tau=1}^t \gamma_\tau y_{t-\tau}\right\}$$

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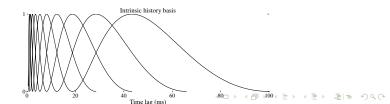
Intrinsic history filter

• Theoretically, model the whole history

$$\lambda_t(t|\theta, y_{t-1}, y_{t-2}, ..., y_1) \propto \exp\left\{\sum_{\tau=1}^t \gamma_\tau y_{t-\tau}\right\}$$

 In practice: use finite history duration with basis functions (a form of regularization: enforce smoothness and reduce parameters)

$$\lambda_t(t|\theta, y_{t-\tau:t-1}) \propto \exp\left\{B \cdot y_{t-\tau:t-1}\right\}$$



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Ensemble history filter

• Ensemble history filters are treated similarly to intrinsic history

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Ensemble history filter

- Ensemble history filters are treated similarly to intrinsic history
- Typically use fewer basis functions than for the intrinsic history, otherwise the number of parameters becomes prohibitive

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- Ensemble history filters are treated similarly to intrinsic history
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- If using regularization, typically each neuron's parameters are penalized as a group (sparse connectivity prior)
- More on intrinsic history and network connectivity next week

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Extrinsic covariates: kinematics

• Truccolo et al. 2005 explore a 2D velocity tuning model based on the x and y components of hand velocity.

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Extrinsic covariates: kinematics

- Truccolo et al. 2005 explore a 2D velocity tuning model based on the x and y components of hand velocity.
- Hatsopoulos et al. 2007 use normalized, extended velocity trajectories "pathlets"

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Maximum likelihood approach to model fitting

• Let y_t be an inhomogeneous Poisson with time varying rate λ_t

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Maximum likelihood approach to model fitting

- Let y_t be an inhomogeneous Poisson with time varying rate λ_t
- The probability of observing k spikes in a time interval Δ is Poisson distributed

$$\Pr(y_t = k) \approx (\Delta \lambda_t)^{y_t} \frac{e^{-\Delta \lambda_t}}{y_t!}$$

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Maximum likelihood approach to model fitting

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$$\Pr(y_t = k) \approx (\Delta \lambda_t)^{y_t} \frac{e^{-\Delta \lambda_t}}{y_t!}$$

• Assuming conditional independence, the probability of observing an entire sequence y_t is

$$\Pr(y|\lambda) = \prod_{t=1}^{T} (\Delta \lambda_t)^{y_t} e^{-\Delta \lambda_t} / y_t!$$

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Minimize the negative log-likelihood

• Fit the model by finding the parameters μ , B that maximize the likelihood $\mathcal{L}(\mu, B|y) = \Pr(y|\mu, B)$ of the observations¹

$$\Pr(y|\mu, B) = \prod_{t=1}^{T} \lambda_t^{y_t} e^{-\lambda_t} / y_t!$$

$$\lambda_t = \exp(\mu + X_t B)$$

¹Note: writing λ_t here instead of $\Delta \lambda_t$, i.e. let $\Delta = 1$. In this case, parameters μ and B will take units of $\Delta = -0$

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Minimize the negative log-likelihood

 Fit the model by finding the parameters μ, B that maximize the likelihood L(μ, B|y) = Pr(y|μ, B) of the observations¹

$$\Pr(y|\mu, B) = \prod_{t=1}^{T} \lambda_t^{y_t} e^{-\lambda_t} / y_t!$$

$$\lambda_t = \exp(\mu + X_t B)$$

• In practice, minimize the negative log-likelihood, which, if Δ is small *s.t.* y_t is always 0 or 1

$$-\ln \mathcal{L}(\mu, B|y) = \sum_{t=1}^{T} [\lambda_t - y_t \ln(\lambda_t)]$$

¹Note: writing λ_t here instead of $\Delta \lambda_t$, i.e. let $\Delta = 1$. In this case, parameters μ and B will take units of $\Delta = \sqrt{2}$ (>)

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Gradient of the negative log-likelihood

- Let $f(\mu,B)$ be the negative log likelihood

$$f(\mu, B) = -\ln \mathcal{L}(\mu, B|y) = \sum_{t=1}^{T} [\lambda_t - y_t \ln(\lambda_t)]$$

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$$f(\mu, B) = -\ln \mathcal{L}(\mu, B|y) = \sum_{t=1}^{T} [\lambda_t - y_t \ln(\lambda_t)]$$

• Substitute our model $\lambda_t = e^{\mu + X_t B}$

$$f(\mu, B) = \sum_{t=1}^{T} [e^{\mu + X_t B} - y_t(\mu + X_t B)]$$

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Gradient of the negative log-likelihood

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• Substitute our model $\lambda_t = e^{\mu + X_t B}$

$$f(\mu, B) = \sum_{t=1}^{T} [e^{\mu + X_t B} - y_t(\mu + X_t B)]$$

• The partial derivatives, w.r.t μ and $(\beta_1, \beta_2, ...) = B$ are:

$$\frac{\partial f}{\partial \mu} = \sum_{t=1}^{T} [e^{\mu + X_t B} - y_t]$$
$$\frac{\partial f}{\partial \beta_i} = \sum_{t=1}^{T} [X_{t,i} e^{\mu + X_t B} - y_t X_{t,i}]$$

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Regularized GLM

- Cross-validation is necessary to assess over-fitting:
 - · data should be separated into training and testing sets

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Regularized GLM

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- For larger number of parameters, the model will overfit:
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Regularized GLM

- Cross-validation is necessary to assess over-fitting:
 - data should be separated into training and testing sets
- For larger number of parameters, the model will overfit:
 - regularization is necessary.
- Regularization can be incorporated by adding a penalty term to the negative log-likelihood function

 $\underset{\mu,B}{\operatorname{argmin}} \left\{ \frac{\operatorname{\mathbf{Penalty}}(B) - \ln \mathcal{L}(\mu, B|y) \right\}$

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Regularized GLM

- Cross-validation is necessary to assess over-fitting:
 - data should be separated into training and testing sets
- For larger number of parameters, the model will overfit:
 - regularization is necessary.
- Regularization can be incorporated by adding a penalty term to the negative log-likelihood function

 $\underset{\mu,B}{\operatorname{argmin}} \left\{ \frac{\operatorname{\mathbf{Penalty}}(B) - \ln \mathcal{L}(\mu, B|y) \right\}$

• Conjugate gradient solvers are useful here

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Regularized GLM: L1 and L2

• L2 penalty: Parameters penalized by their squared magnitudes



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· Equivalent to a Gaussian prior on parameters

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- · Can be solved with gradient descent

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• L2 penalty: Parameters penalized by their squared magnitudes



- Equivalent to a Gaussian prior on parameters
- · Can be solved with gradient descent
- L1 penalty: Parameters penalized by their absolute magnitude

$$\alpha \sum_{i=1}^{N} |\beta_i|$$

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$$\alpha \sum_{i=1}^{N} |\beta_i|$$

- Promotes $\beta_i = 0$, useful for finding sparse solutions

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$$\alpha \sum_{i=1}^{N} |\beta_i|$$

- Promotes $\beta_i = 0$, useful for finding sparse solutions
- Discontinuous gradient at $\beta_i = 0$ percludes gradient descent.
- Can use coordinate descent (although we ran into convergence issues?)

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Regularized GLM: L1 approximation and L0

• $\sqrt{x^2+\epsilon}$: Smooth approximation of L1 penalty that is suitable for gradient descent^2

$$\alpha \sum_{i=1}^N \sqrt{\beta_i^2 + \varepsilon}$$

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- ϵ is chosen to be small, but strictly positive

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- + ϵ is chosen to be small, but strictly positive
- L0 penalty: Constant penalty if a parameter is nonzero

$$\alpha \sum_{i=1}^{N} \delta(\beta_i \neq 0)$$

²Called "Charbonnier penalty" in the computer vision literature (Charbonnier et al.(1994) = > + + = > = = - > <

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• Computationally infeasible

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Regularized GLM: L1 approximation and L0

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- Computationally infeasible
- Greedy algorithms are a good (the best polynomial time?) approximation

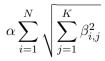
²Called "Charbonnier penalty" in the computer vision literature (Charbonnier et al. (1994) $\in \mathbb{R} \times \mathbb{R} \times \mathbb{R} \times \mathbb{R} \times \mathbb{R}$

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Regularized GLM: Group lasso

• Concept: penalize groups of parameters with the L1 norm



³Roth, Volker, and Bernd Fischer. "The group-lasso for generalized linear models: uniqueness of solutions and efficient algorithms." Proceedings of the 25th international conference on Machine learning. ACM; 2008: • 🚊 😑 🗠 🔍

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• Concept: penalize groups of parameters with the L1 norm

$$\alpha \sum_{i=1}^{N} \sqrt{\sum_{j=1}^{K} \beta_{i,j}^2}$$

• Useful for penalizing ensemble filters as a group for each neuron

³Roth, Volker, and Bernd Fischer. "The group-lasso for generalized linear models: uniqueness of solutions and efficient algorithms." Proceedings of the 25th international conference on Machine learning. ACM; 2008; * 🚊 😑 🗠 🔍 (>

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- Useful for penalizing ensemble filters as a group for each neuron
- Derivative is undefined at $\sqrt{\sum_{j=1}^{K} \beta_{i,j}^2} = 0$
- Roth & Fischer 2008³ discuss an efficient fitting procedure

³Roth, Volker, and Bernd Fischer. "The group-lasso for generalized linear models: uniqueness of solutions and efficient algorithms." Proceedings of the 25th international conference on Machine learning. ACM; 2008; * 🚊 😑 🗠 🔍 (>

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Two-layer crossvalidation

- The strength of regularization $\boldsymbol{\alpha}$ is another free parameter

• Can be fixed in advance or...

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Two-layer crossvalidation

- The strength of regularization α is another free parameter
 - Can be fixed in advance or...
 - Two-level crossvalidation can be used to estimate α from the data

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Two-layer crossvalidation

- The strength of regularization $\boldsymbol{\alpha}$ is another free parameter
 - Can be fixed in advance or...
 - Two-level crossvalidation can be used to estimate α from the data
- K-fold two-layer crossvalidation results in K^2 parameter fitting steps
 - This can get slow on large datasets

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Regularization paths

- Regularization paths:
 - When evaluating a range of regularization parameters e.g. $\alpha = (0, 0.1, 1, 20) \text{,}$
 - Carry over the model weights μ and β .
 - That is, for example, start the parameter tuning for $\alpha=1$ with the μ,B returned from $\alpha=0.1$

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Initialization with closed form solution

• Fitting many models using a large amount of data can take a very long time

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Initialization with closed form solution

- Fitting many models using a large amount of data can take a very long time
- Using an approximation to estimate a good starting location can give some speedup
- Ramirez, A. D., & Paninski, L. (2014). Fast inference in generalized linear models via expected log-likelihoods. Journal of computational neuroscience, 36(2), 215-234.

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$$\lambda(t \mid H(t)) = \lim_{\Delta \to 0} \frac{P[N(t + \Delta) - N(t) = 1 \mid H(t)]}{\Delta}$$
(1)

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$$P(N_{1:K} \mid \theta) = \prod_{k=1}^{K} \left[\lambda(t_k \mid \theta, H_k) \Delta \right]^{\Delta N_k} \left[1 - \lambda(t_k \mid \theta, H_k) \Delta \right]^{1 - \Delta N_k} + o(\Delta^J) \quad (2)$$

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$$P(N_{1:K} \mid \theta) = \exp\left\{\sum_{k=1}^{K} \log\left[\lambda(t_k \mid \theta, H_k)\Delta\right]\Delta N_k - \sum_{k=1}^{K} \lambda(t_k \mid \theta, H_k)\Delta\right\} + o(\Delta^J) \quad (3)$$

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$$\log \lambda(t_k \mid \theta, H_k) = \sum_{i=1}^q \theta_i g_i [v_i(t_k + \tau)]$$
(4)

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$$\log\left\{\left[1-\lambda(t_k \mid \theta, H_k)\Delta\right]^{-1}\left[\lambda(t_k \mid \theta, H_k)\Delta\right]\right\} = \sum_{i=1}^{q} \theta_i g_i [v_i(t_k + \tau)]$$
(5)

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$\lambda(t_k \mid N_{1:k}^{1:C}, \mathbf{x}_{k+\tau}, \theta) = \lambda_l(t_k \mid N_{1:k}, \theta_l) \lambda_E(t_k \mid N_{1:k}^{1:C}, \theta_E) \lambda_X(t_k \mid \mathbf{x}_{k+\tau}, \theta_X)$ (6)

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$$\lambda_{I}(t_{k} \mid N_{1:k}, \theta_{I}) = \exp\left\{\gamma_{0} + \sum_{n=1}^{Q} \gamma_{n} \Delta N_{k-n}\right\}$$
(7)

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$$\lambda_E(t_k \mid N_{1:k}^{1:C}, \theta_E) = \exp\left\{\beta_0 + \sum_c \sum_{r=1}^R \beta_r^c \Delta N_{k-r}^c\right\}$$
(8)

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$$\lambda_{E}(t_{k} \mid N_{1:k}^{1:C}, \theta_{E}) = \exp\left\{\beta_{0} + \sum_{c} \sum_{r=1}^{R} \beta_{r}^{c} \left(N_{k-(r-1)W}^{c} - N_{k-rW}^{c}\right)\right\}$$
(9)

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$\lambda_X(t_k \mid \mathbf{x}_{k+\tau}, \theta_X) = \exp\{\alpha_0 + |V_{k+\tau}| [\alpha_1 \cos(\phi_{k+\tau}) + \alpha_2 \sin(\phi_{k+\tau})]\} \quad (10)$

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$$\lambda(t_k \mid N_{1:k}^{1:C}, \mathbf{x}_{k+\tau}, \theta) = \exp\{\mu + \sum_{n=1}^{Q} \gamma_n \Delta N_{k-n}\}$$

+
$$\sum_{c} \sum_{r=1}^{R} \beta_{r}^{c} \Delta N_{k-r}^{c} + |V_{k+\tau}| [\alpha_{1} \cos{(\phi_{k+\tau})} + \alpha_{2} \sin{(\phi_{k+\tau})}] \}$$
 (11)

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$$\lambda(t_k \mid \theta, H_k) = \frac{p(t_e \mid \theta, H_k)}{1 - \int_{u_{N_{k-1}}}^{t_k} p[t \mid \theta, H(t)] \mathrm{d}t}$$
(12)

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$$s(t_{k} | \mathbf{x}_{t+\tau}, N_{1:k}^{1:C}, \theta_{X}, \theta_{E}) = \exp \left\{ \mu + \sum_{c} \sum_{r=1}^{R} \beta_{r}^{c} (N_{k-(r-1)W}^{c} - N_{k-rW}^{c}) + |V_{k+\tau}| [\alpha_{1} \cos (\phi_{k+\tau}) + \alpha_{2} \sin (\phi_{k+\tau})] \right\}$$
(13)

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$$z_j = 1 - \exp\left\{-\int_{u_j}^{u_{j+1}} \lambda(t \mid H(t), \,\hat{\theta}) \mathrm{d}t\right\}$$
(14)

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$$M(t_k) = \sum_{i=k-B}^{k} \Delta N_i - \int_{t_{k-B}}^{t_k} \lambda(t \mid H(t), \hat{\theta}) dt$$
(15)

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$$AIC(q) = -2\log L(\hat{\theta} \mid H_k) + 2q \tag{16}$$

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$$\mathbf{x}_{k+\tau} = \boldsymbol{\mu}_{\mathbf{x}} + \mathbf{F} \mathbf{x}_{k+\tau-1} + \boldsymbol{\varepsilon}_{k+\tau} \tag{17}$$

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$$\mathbf{x}_{k+\tau|k+\tau-1} = \boldsymbol{\mu}_{\mathbf{x}} + \mathbf{F} \mathbf{x}_{k+\tau-1|k+\tau-1}$$
(18)

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$$\mathbf{W}_{k+\tau|k+\tau-1} = \mathbf{F}\mathbf{W}_{k+\tau-1|k+\tau-1}\mathbf{F} + \mathbf{W}_{\varepsilon}$$
(19)

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$$\mathbf{W}_{k+\tau|k+\tau} = \left[\mathbf{W}_{k+\tau|k+\tau-1}^{-1} + \sum_{c=1}^{C} \left[\nabla \log \lambda^{c} \left(t_{k} \mid N_{1:k}^{c}, \mathbf{x}_{k+\tau|k+\tau-1}, \hat{\theta}^{c} \right) \right] \right]$$

$$\lambda^{c} \left(t_{k} \mid N_{1:k}^{c}, \mathbf{x}_{k+\tau|k+\tau-1}, \hat{\theta}^{c} \right) \Delta \left[\nabla \log \lambda^{c} \left(t_{k} \mid N_{1:k}^{c}, \mathbf{x}_{k+\tau|k+\tau-1}, \hat{\theta}^{c} \right) \right] \right]$$

$$- \sum_{c=1}^{C} \nabla^{2} \log \lambda^{c} \left(t_{k} \mid N_{1:k}^{c}, \mathbf{x}_{k+\tau|k+\tau-1}, \hat{\theta}^{c} \right) \left[\Delta N_{1:k}^{c} - \lambda^{c} \left(t_{k} \mid N_{1:k}^{c}, \mathbf{x}_{k+\tau|k+\tau-1}, \hat{\theta}^{c} \right) \Delta \right] \right]^{-1} \quad (20)$$

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$$\mathbf{x}_{k+\tau|k+\tau} = \mathbf{x}_{k+\tau|k+\tau-1} + \mathbf{W}_{k+\tau|k+\tau}$$

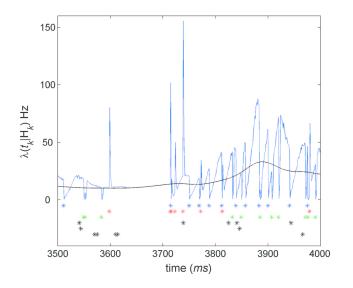
$$\times \sum_{c=1}^{C} \nabla \log \lambda^{c}(t_{k} \mid N_{1:k}^{c}, \mathbf{x}_{k+\tau|k+\tau-1}, \hat{\theta}^{c}) \left[\Delta N_{1:k}^{c} - \lambda^{c}(t_{k} \mid N_{1:k}^{c}, \mathbf{x}_{k+\tau|k+\tau-1}, \hat{\theta}^{c})\Delta\right]$$
(21)

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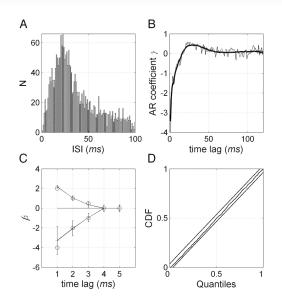
$$(\mathbf{x}_{k+\tau} - \mathbf{x}_{k+\tau|k+\tau})' \mathbf{W}_{k+\tau|k+\tau}^{-1} (\mathbf{x}_{k+\tau} - \mathbf{x}_{k+\tau|k+\tau}) \le \chi_{0.95}^2(m)$$
(22)

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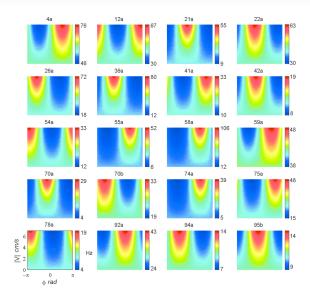
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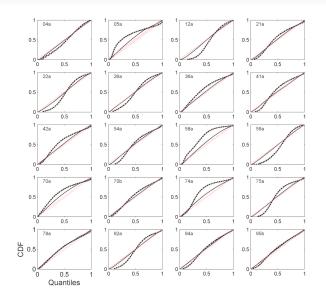
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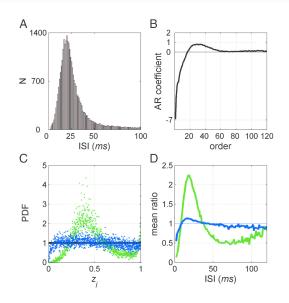
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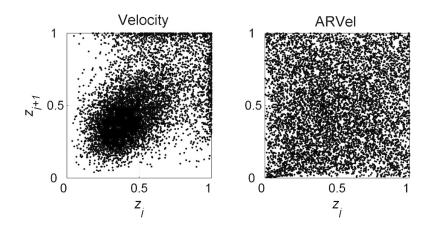
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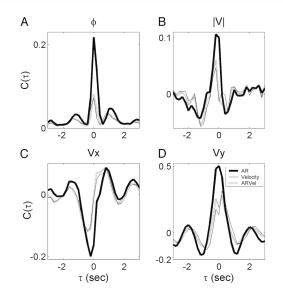
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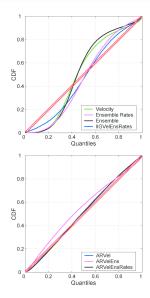
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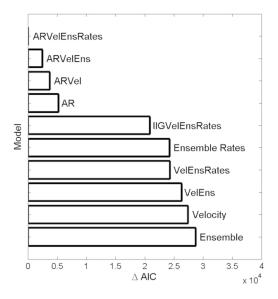
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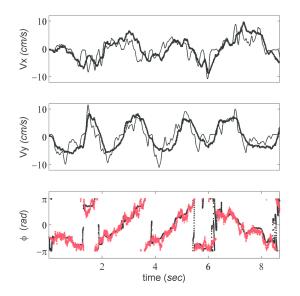


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$$A_{k} = \{ \text{spike in } (t_{k-1}, t_{k}] \mid H_{k} \}$$

$$E_{k} = \{A_{k}\}^{\Delta N_{k}} \{A_{k}^{c}\}^{1-\Delta N_{k}}$$

$$H_{k} = \left\{ \bigcap_{j=1}^{k-1} E_{j} \right\}$$
(A1)

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$$P(N_{1:K}) = P(u_j \in (t_{k_j-1}, t_{k_j}], j = 1, \dots, J, \cap N(T) = J)$$
$$= \prod_{k=1}^{K} P(A_k)^{\Delta N_k} P(A_k^c)^{1-\Delta N_k}$$
(A2)

GLM point process for spike trains

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Truccolo et al. 2005 figures and equations

$$P(A_k) = \lambda(t_k \mid H_k)\Delta + o(\Delta)$$

$$P(A_k^c) = 1 - \lambda(t_k \mid H_k)\Delta + o(\Delta)$$
(A3)

1.4.3

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

$$P(N_{1:K}) = \prod_{k=1}^{K} \left[\lambda(t_k \mid H_k) \Delta \right]^{\Delta N_k} \left[1 - \lambda(t_k \mid H_k) \Delta \right]^{1 - \Delta N_k} + o(\Delta^J)$$
(A4)

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

$$P(N_{1:K}) = \prod_{k=1}^{K} \left[\lambda(t_k \mid H_k) \Delta \right]^{\Delta N_k} \left[1 - \lambda(t_k \mid H_k) \Delta \right]^{-\Delta N_k} \prod_{k=1}^{K} \left[1 - \lambda(t_k \mid H_k) \Delta \right] + o(\Delta^J)$$
$$= \prod_{k=1}^{K} \left[\frac{\lambda(t_k \mid H_k) \Delta}{1 - \lambda(t_k \mid H_k) \Delta} \right]^{\Delta N_k} \prod_{k=1}^{K} \exp\{-\lambda(t_k \mid H_k) \Delta\} + o(\Delta^J)$$
$$= \exp\left\{ \sum_{k=1}^{K} \log \left[\lambda(t_k \mid H_k) \Delta \right] \Delta N_k - \sum_{k=1}^{K} \lambda(t_k \mid H_k) \Delta \right\} + o(\Delta^J)$$
(A5)

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

$$P(N_{0:T}) = \lim_{\Delta \to 0} \frac{\exp\left\{\sum_{k=1}^{K} \log\left[\lambda(t_k \mid H_k)\Delta\right]\Delta N_k - \sum_{k=1}^{K} \lambda(t_k \mid H_k)\Delta\right\} + o(\Delta^J)}{\Delta^J}$$

$$= \lim_{\Delta \to 0} \frac{\exp\left\{\sum_{k=1}^{K} \log \lambda(t_k \mid H_k) \Delta N_k - \sum_{k=1}^{K} \lambda(t_k \mid H_k) \Delta\right\} \Delta^J + o(\Delta^J)}{\Delta^J}$$

$$= \exp\left\{\int_{0}^{T} \log \lambda(t \mid H(t)) dN(t) - \int_{0}^{T} \lambda(t \mid H(t)) dt\right\}$$
(A6)

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

$$f(\mathbf{y} \mid \mathbf{\Theta}, \phi) = \exp\{[\mathbf{y}\mathbf{\Theta} - b(\mathbf{\Theta})]/a(\phi) + c(\mathbf{y}, \phi)\}$$
(A7)

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

$$f(\mathbf{y} \mid \mathbf{\Theta}, \phi) \propto \exp\left\{\sum_{k=1}^{K} y_k \log \lambda_k - \sum_{k=1}^{K} \lambda_k\right\}$$
 (A8)

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

$$f(\mathbf{y} \mid \mathbf{\Theta}, \phi) = \prod_{k=1}^{K} [P_k]^{y_k} [1 - P_k]^{1 - y_k}$$
(A9)

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

$$\log\left[\frac{P(A_k \mid \theta, H_k)}{1 - P(A_k \mid \theta, H_k)}\right] \approx \log\left[\lambda(t_k \mid \theta, H_k)\Delta\right]$$
(A10)

GLM point process for spike trains

Truccolo et al. 2005 figures and equations

END

Generalized linear model: nonlinearity on the response variable

F(y) = XB



Regularized GLM: Elastic net

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• Concept: dynamic trade-off between L1 and L2

Regularized GLM: Elastic net

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- Concept: dynamic trade-off between L1 and L2
 - Code crashed pretty hard.