Derivatives of Gaussian KL-Divergence for some parameterizations of the posterior covariance for variational Gaussian-process inference

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March 25, 2020

These notes provide the derivatives of the KL-divergence $D_{KL}[Q(\mathbf{z})||P(\mathbf{z})]$ between two multivariate Gaussian distributions $Q(\mathbf{z})$ and $P(\mathbf{z})$ with respect to a few parameterizations θ of the covariance matrix $\Sigma(\theta)$ of Q. This is useful for variational Gaussian process inference, where clever parameterizations of the posterior covariance are required to make the problem tractable. Tables for differentiating matrix-valued functions can be found in The Matrix Cookbook.

Consider two multivariate Gaussian distributions $Q(\mathbf{z}) = \mathcal{N}(\boldsymbol{\mu}_q, \boldsymbol{\Sigma}(\theta))$ and $P(\mathbf{z}) = \mathcal{N}(\boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0 = \boldsymbol{\Lambda}^{-1})$ with dimension *L*. The KL divergence $D_{\text{KL}}[Q(\mathbf{z})||P(\mathbf{z})]$ has the closed form

$$\mathcal{D} := D_{\mathrm{KL}} \left[Q(\mathbf{z}) \| \operatorname{Pr}(\mathbf{z}) \right]$$

= $\frac{1}{2} \left\{ (\boldsymbol{\mu}_0 - \boldsymbol{\mu}_q)^{\mathsf{T}} \boldsymbol{\Lambda} (\boldsymbol{\mu}_0 - \boldsymbol{\mu}_q) + \operatorname{tr} (\boldsymbol{\Lambda} \boldsymbol{\Sigma}) - \ln |\boldsymbol{\Sigma}| - \ln |\boldsymbol{\Lambda}| \right\} + \text{constant.}$ (1)

In variational Bayesian inference, we minimize \mathcal{D} while maximizing the expected log-probability of some observations with respect to $Q(\mathbf{z})$. Closed-form derivatives of \mathcal{D} in terms of the parameters of Q are useful for manually optimizing code for larger problems. The derivatives of \mathcal{D} in terms of $\boldsymbol{\mu}_q$ are straightforward: $\partial_{\boldsymbol{\mu}_q} \mathcal{D} = \Lambda(\boldsymbol{\mu}_q - \boldsymbol{\mu}_z)$ and $H_{\boldsymbol{\mu}_q} \mathcal{D} = \Lambda$. In these notes, we explore derivatives of \mathcal{D} with respect to a few different parameterizations (" θ ") of $\Sigma(\theta)$.

We evaluate the following parameterizations for Σ : 1. Optimizing the full Σ directly 2. $\Sigma \approx \mathbf{X}\mathbf{X}^{\top}$ 3. $\Sigma \approx \mathbf{A}^{\top} \operatorname{diag}[\mathbf{v}]\mathbf{A}$ 4. $\Sigma \approx [\mathbf{\Lambda} + \operatorname{diag}[\mathbf{p}]]^{-1}$ 5. $\mathbf{F}^{\top}\mathbf{Q}\mathbf{Q}^{\top}\mathbf{F}$, where $\mathbf{Q} \in \mathbb{R}^{K \times K}$, K < L and $\mathbf{F} \in \mathbb{R}^{K \times L}$, $\mathbf{F}\mathbf{F}^{\top} = \mathbf{I}$.

0.1 Σ

We first obtain gradients of \mathcal{D} in Σ (assuming Σ is full-rank). These can be used to derive gradients in θ for some parameterizations $\Sigma(\theta)$ using the chain rule. The gradient of \mathcal{D} in Σ can be obtained using identities (57) and (100) in The Matrix Cookbook:

$$\partial_{\Sigma} \mathcal{D} = \partial_{\Sigma} \left\{ \operatorname{tr} \left(\Lambda \Sigma \right) - \ln |\Sigma| \right\} = \frac{1}{2} \left(\Lambda - \Sigma^{-1} \right).$$
⁽²⁾

The Hessian in Σ is a fourth-order tensor. It's simpler to express the Hessian in terms of a Hessian-vector product, which can be used with Krylov subspace solvers to efficiently compute the update in Newton's method. Considering an *L*×*L* matrix **M**, the Hessian-vector product is given by

$$[\mathbf{H}_{\Sigma}\mathcal{D}]\mathbf{M} = \partial_{\Sigma} \langle \partial_{\Sigma}\mathcal{D}, \mathbf{M} \rangle = \partial_{\Sigma} \operatorname{tr} \left[(\partial_{\Sigma}\mathcal{D})^{\mathsf{T}} \mathbf{M} \right],$$
(3)

where $\langle \cdot, \cdot \rangle$ denotes the scalar (Frobenius) product. This is given by identity (124) in the Matrix Cookbook:

$$\partial_{\Sigma} \operatorname{tr} \left[\frac{1}{2} \left(\boldsymbol{\Lambda} - \boldsymbol{\Sigma}^{-1} \right)^{\top} \mathbf{M} \right] = -\frac{1}{2} \partial_{\Sigma} \operatorname{tr} \left[\boldsymbol{\Sigma}^{-1} \mathbf{M} \right] = \frac{1}{2} \boldsymbol{\Sigma}^{-1} \mathbf{M}^{\top} \boldsymbol{\Sigma}^{-1}.$$
(4)

0.2 $\Sigma \approx X X^{\top}$

We consider an approximate posterior covariance of the form

$$\Sigma \approx \mathbf{X}\mathbf{X}^{\mathsf{T}}, \quad \mathbf{X} \in \mathbb{R}^{L \times K}$$
 (5)

where **X** is a rank-K < L matrix with *L* rows and *K* columns.

Since X is not full rank, the log-determinant $\ln |\Sigma| = \ln |XX^{\top}|$ in (1) diverges, due to the zero eigenvalues in the null space of X. However, since this null-space is not being optimized, it does not affect our gradient. It is sufficient to replace the log-determinant with that of the reduced-rank representation, $\ln |X^{\top}X|$. Identity (55) in The Matrix Cookbook provides the derivative of this, $\partial_X \ln |X^{\top}X| = 2X^{+\top}$, where $(\cdot)^+$ is the pseudoinverse. Combined with identity (112), this gives the following gradient of $\mathcal{D}(X)$:

$$\partial_{\mathbf{X}} \mathcal{D} = \partial_{\mathbf{X}} \frac{1}{2} \left\{ \operatorname{tr} \left[\mathbf{\Lambda} \mathbf{X} \mathbf{X}^{\top} \right] - \ln |\mathbf{X}^{\top} \mathbf{X}| \right\} = \mathbf{\Lambda} \mathbf{X} - \mathbf{X}^{+\top}.$$
(6)

The Hessian-vector product requires the derivative of ∂_X tr [X⁺M]:

$$\partial_{\mathbf{X}} \langle \partial_{\mathbf{\Sigma}} \mathcal{D}, \mathbf{M} \rangle = \partial_{\mathbf{X}} \operatorname{tr} \left[\left(\mathbf{\Lambda} \mathbf{X} - \mathbf{X}^{+\top} \right)^{\top} \mathbf{M} \right] = \partial_{\mathbf{X}} \operatorname{tr} \left[\mathbf{\Lambda} \mathbf{X} \mathbf{M} \right] - \partial_{\mathbf{X}} \operatorname{tr} \left[\mathbf{X}^{+} \mathbf{M} \right].$$
(7)

Goulob and Pereya (1972) Eq. 4.12 gives the derivative of a fixed-rank pseudoinverse:

$$\partial X^{+} = -X^{+}(\partial X)X^{+} + X^{+}X^{+\top}(\partial X)^{\top}(1 - XX^{+}) + (1 - X^{+}X)(\partial X)^{\top}X^{+\top}X^{+}$$
(8)

Since X is $N \times K$ with rank K, X⁺X is full-rank. Therefore X⁺X = I_k and the final term in (??) vanishes. The derivative of the pseudoinverse can now be written as:

$$\partial \mathbf{X}^{+} = -\mathbf{X}^{+}(\partial \mathbf{X})\mathbf{X}^{+} + \mathbf{X}^{+}\mathbf{X}^{+\top}(\partial \mathbf{X})^{\top}(\mathbf{I}_{n} - \mathbf{X}\mathbf{X}^{+})$$
(9)

Since the derivative of a trace of a matrix-valued function is just the (transpose) of the scalar derivative,

$$\partial_{\mathbf{X}} \operatorname{tr} \left[\mathbf{X}^{+} \mathbf{M} \right] = \left\{ -\mathbf{X}^{+} \mathbf{M} \mathbf{X}^{+} + \mathbf{X}^{+} \mathbf{X}^{+\top} \mathbf{M}^{\top} (\mathbf{I}_{n} - \mathbf{X} \mathbf{X}^{+}) \right\}^{\top} = -\mathbf{X}^{+\top} \mathbf{M}^{\top} \mathbf{X}^{+\top} + (\mathbf{I} - \mathbf{X}^{+\top} \mathbf{X}^{\top}) \mathbf{M} \mathbf{X}^{+} \mathbf{X}^{+\top}.$$
(10)

Overall, we obtain the following Hessian-vector product:

$$\partial_{\mathbf{X}} \langle \partial_{\Sigma} \mathcal{D}, \mathbf{M} \rangle = \mathbf{\Lambda} \mathbf{M}^{\top} + \mathbf{X}^{+\top} \mathbf{M}^{\top} \mathbf{X}^{+\top} - (\mathbf{I} - \mathbf{X}^{+\top} \mathbf{X}^{\top}) \mathbf{M} \mathbf{X}^{+} \mathbf{X}^{+\top}$$
(11)

0.2.1 $\Sigma = XX^{\top}$ when X is full-rank

Equations (6) and (11) are also valid if X is a rank-*L* triangular (Choleskey) factorization of Σ . In this case the pseudoinverse can be replaced by the full inverse, and various terms simplify:

$$\partial_{\mathbf{X}} \mathcal{D} = \mathbf{\Lambda} \mathbf{X} - \mathbf{X}^{-\top}$$
$$\partial_{\mathbf{X}} \langle \partial_{\mathbf{x}} \mathcal{D}, \mathbf{M} \rangle = \mathbf{\Lambda} \mathbf{M}^{\top} + \mathbf{X}^{-\top} \mathbf{M}^{\top} \mathbf{X}^{-\top}$$
(12)

0.3 $\Sigma = \mathbf{A}^{\top} \operatorname{diag}[\mathbf{v}]\mathbf{A}$

Let $\Sigma = \mathbf{A}^{\top} \operatorname{diag}[\mathbf{v}]\mathbf{A}$, where \mathbf{A} is fixed and $\mathbf{v} \in \mathbb{R}^{L}$ are free parameters. Define $\operatorname{diag}[\cdot]$ as an operator that constructs a diagonal matrix from a vector, or extracts the main diagonal from a matrix if its argument is a matrix. The gradient of \mathcal{D} in \mathbf{v} is:

$$\partial_{\mathbf{X}} \mathcal{D} = \partial_{\mathbf{X}} \frac{1}{2} \left\{ \operatorname{tr} \left[\mathbf{\Lambda} \mathbf{\Lambda}^{\top} \operatorname{diag}[\mathbf{v}] \mathbf{A} \right] - \ln |\mathbf{A}^{\top} \operatorname{diag}[\mathbf{v}] \mathbf{A} | \right\} \\ = \frac{1}{2} \left\{ \operatorname{diag}[\mathbf{A} \mathbf{\Lambda} \mathbf{A}^{\top}] - \frac{1}{\mathbf{v}} \right\}$$
(13)

The hessian in **v** is a matrix in this case:

$$H_{\mathbf{v}}\mathcal{D} = \frac{1}{2}\operatorname{diag}\left[\frac{1}{\mathbf{v}^{2}}\right].$$
(14)

This parameterization is useful for spatiotemporal inference problems, where the matrix A represents a fixed convolution which can be evaluated using the Fast Fourier Transform (FFT).

0.4 Inverse-diagonal approximation

Let $\Sigma^{-1} = \Lambda + \text{diag}[\mathbf{p}]$. To obtain the gradient in \mathbf{p} , combine the derivatives $\partial_{\Sigma} \mathcal{D}$ (Eq. (2)) and $\partial_{\mathbf{p}} \Sigma$ using the chain rule. If $\{(\Sigma) \text{ is a function of } \Sigma, \text{ and } \Sigma(\theta_i) \text{ is a function of a parameter } \theta_i$, then the chain rule is (The Matrix Cookbook; Eq. 136):

$$\partial_{\theta_i} \{ = \left\langle \partial_{\Sigma} \{, \partial_{\theta_i} \Sigma \right\rangle = \sum_{kl} (\partial_{\Sigma_{kl}} \{) (\partial_{\theta_i} \Sigma_{kl})$$
(15)

From (2) we have $\partial_{\Sigma} \mathcal{D} = \frac{1}{2} (\Lambda - \Sigma^{-1})$; Since $\Sigma^{-1} = \Lambda + \text{diag} [\mathbf{p}]$, this simplifies to:

$$\partial_{\Sigma} \mathcal{D} = \frac{1}{2} \left(\mathbf{\Lambda} - \mathbf{\Sigma}^{-1} \right)$$

= $\frac{1}{2} \left(\mathbf{\Lambda} - \mathbf{\Lambda} - \text{diag} \left[\mathbf{p} \right] \right)$
= $-\frac{1}{2} \text{diag} \left[\mathbf{p} \right]$ (16)

We also need $\partial_{\mathbf{p}_i} \Sigma$. Let $\mathbf{Y} = \Sigma^{-1}$. The derivative $\partial \mathbf{Y}^{-1}$ is given as identity (59) in The Matrix Cookbook as $\partial \mathbf{Y}^{-1} = -\mathbf{Y}^{-1}(\partial \mathbf{Y})\mathbf{Y}^{-1}$. Using this, we can obtain $\partial_{\mathbf{p}_i} \Sigma$:

$$\partial_{\mathbf{p}_{i}}\Sigma = \partial_{\mathbf{p}_{i}}\mathbf{Y}^{-1} = -\mathbf{Y}^{-1} \left(\partial_{\mathbf{p}_{i}}\mathbf{Y}\right)\mathbf{Y}^{-1} = -\Sigma \left(\partial_{\mathbf{p}_{i}}\Sigma^{-1}\right)\Sigma$$
$$= -\Sigma \partial_{\mathbf{p}_{i}}\left[\mathbf{\Lambda} + \operatorname{diag}[\mathbf{p}_{i}]\right]\Sigma = -\Sigma \mathbf{J}_{ii}\Sigma$$
$$= -\boldsymbol{\sigma}_{i}\boldsymbol{\sigma}_{i}^{\mathsf{T}}$$
(17)

where σ_i is the *i*th row of Σ and J_{ii} is a matrix which is zero everywhere, except for at index (*i*, *i*), where it is 1.

Applying (15) to (16) and (17) for a particular element \mathbf{p}_i gives:

$$\partial_{\mathbf{p}_{i}} \mathcal{D} = \sum_{kl} [\partial_{\Sigma_{kl}} \mathcal{D}] [\partial_{\mathbf{p}_{i}} \Sigma_{kl}] = \sum_{kl} \left\{ -\frac{1}{2} \operatorname{diag} [\mathbf{p}] \right\}_{kl} \left\{ -\sigma_{i} \sigma_{i}^{\top} \right\}_{kl}$$
$$= \frac{1}{2} \sum_{kl} \delta_{k=l} \mathbf{p}_{k} \sigma_{ik} \sigma_{il} = \frac{1}{2} \sum_{k} \mathbf{p}_{k} \sigma_{ik} \sigma_{ik} = \frac{1}{2} \sum_{k} \mathbf{p}_{k} \sigma_{ik}^{2}$$
$$= \frac{1}{2} \mathbf{p} \sigma_{i}^{\circ 2}$$
(18)

where $(\cdot)^{\circ 2}$ denotes the element-wise square of a vector or matrix. In matrix notation, this is:

$$\partial_{\mathbf{p}} \mathcal{D} = \frac{1}{2} \mathbf{p} \Sigma^{\circ 2} = \frac{1}{2} \operatorname{diag} \left[\Sigma \operatorname{diag} \left[\mathbf{p} \right] \Sigma \right], \tag{19}$$

The Hessian-vector product is cumbersome, since each term in the expression Σ (diag [**p**]) Σ depends on **p**. In the case of the log-linear Poisson GLM, the gradient (??) simplifies further and optimization becomes tractable. We will explore this further in later notes.

This parameterization resembles the closed-form covariance update for a linear, Gaussian model, where 1/p is a vector of measurement noise variances. It is also a useful parameterization for variational Bayesian solutions for non-conjugate Generalized Linear Models (GLMs), where **p** becomes a free parameter to be estimated.

$\mathbf{0.5} \quad \boldsymbol{\Sigma} = \mathbf{F}^{\top} \mathbf{Q} \mathbf{Q}^{\top} \mathbf{F}$

Let $\Sigma = \mathbf{F}^{\mathsf{T}} \mathbf{Q} \mathbf{Q}^{\mathsf{T}} \mathbf{F}$, where $\mathbf{Q} \in \mathbb{R}^{K \times K}$; K < L is the free parameter and $\mathbf{F} \in \mathbb{R}^{K \times L}$ is a fixed transformation. If \mathbf{Q} is a lower-triangular matrix, then this approximation involves optimizing K(K + 1)/2 parameters.

Since the trace is invariant under cyclic permutation, tr $[\Lambda F^{\top}QQ^{\top}F] = \text{tr} [F\Lambda F^{\top}QQ^{\top}]$. The derivatives have the same form as (12) with $\tilde{\Lambda} = F\Lambda F^{\top}$:

$$\partial_{\mathbf{Q}}\mathcal{D} = \tilde{\Lambda}\mathbf{Q} - \mathbf{Q}^{-\top}$$

= $\mathbf{F}\Lambda\mathbf{F}^{\top}\mathbf{Q} - \mathbf{Q}^{-\top}$
 $\partial_{\mathbf{Q}}\left\langle\partial_{\mathbf{Q}}\mathcal{D},\mathbf{M}\right\rangle = \tilde{\Lambda}\mathbf{M}^{\top} + \mathbf{Q}^{-\top}\mathbf{M}^{\top}\mathbf{Q}^{-\top}$
= $\mathbf{F}\Lambda\mathbf{F}^{\top}\mathbf{M}^{\top} + \mathbf{Q}^{-\top}\mathbf{M}^{\top}\mathbf{Q}^{-\top}$ (20)

This form is convenient for spatiotemporal inference problems that are sparse in frequency space. In this application, **F** corresponds a (unitary) Fourier transform with all by *K* of the resulting frequency components discarded. The product of **F** with a vector **v** can be computed in $O[L \log(L)]$ time using the Fast Fourier Transform (FFT). Alternatively, if $K \leq O(\log(L))$, it is faster to simply multiply **Fv** directly. Furthermore, if **F** is semi-orthogonal (**FF**^{\top} = **I**), then calculation of **F**^{\top}**Q** can be re-used (for example diag[Σ] = [(**F**^{\top}**Q**)^{°2}]^{\top}**1**).

0.6 Conclusion

These notes provide the gradients and Hessian-vector products for four simplified parameterizations of the posterior covariance matrix for variational Gaussian process inference. If combined with the gradients

and Hessian-vector products for the expected log-likelihood, these expressions can be used with Krylov-subspace solvers to compute the Newton-Raphson update to optimize Σ .

We evaluated the following parameterizations for Σ : 1. Σ :

$$\partial = \frac{1}{2} \left(\mathbf{\Lambda} - \boldsymbol{\Sigma}^{-1} \right)$$

$$\partial \left\langle \partial, \mathbf{M} \right\rangle = \frac{1}{2} \boldsymbol{\Sigma}^{-1} \mathbf{M}^{\mathsf{T}} \boldsymbol{\Sigma}^{-1}$$
(21)

2. $\Sigma \approx \mathbf{X}\mathbf{X}^{\top}$:

$$\partial = \Lambda \mathbf{X} - \mathbf{X}^{+\top}.$$

$$\partial \langle \partial, \mathbf{M} \rangle = \Lambda \mathbf{M}^{\top} + \mathbf{X}^{+\top} \mathbf{M}^{\top} \mathbf{X}^{+\top} - (\mathbf{I} - \mathbf{X}^{+\top} \mathbf{X}^{\top}) \mathbf{M} \mathbf{X}^{+} \mathbf{X}^{+\top}$$
(22)

3. $\Sigma \approx \mathbf{A}^{\top} \operatorname{diag}[\mathbf{v}]\mathbf{A}$:

$$\partial = \frac{1}{2} \left\{ \operatorname{diag} \left[\mathbf{A} \mathbf{A} \mathbf{A}^{\mathsf{T}} \right] - \frac{1}{\mathbf{v}} \right\}$$

$$\partial \left\langle \partial, \mathbf{u} \right\rangle = \frac{1}{2} \left[\frac{1}{\mathbf{v}^2} \right]^{\mathsf{T}} \mathbf{u}$$
(23)

4. $\Sigma \approx [\Lambda + \text{diag}[\mathbf{p}]]^{-1}$:

$$\partial = \frac{1}{2} \mathbf{p} \Sigma^{\circ 2} = \frac{1}{2} \operatorname{diag} \left[\Sigma \operatorname{diag} \left[\mathbf{p} \right] \Sigma \right], \tag{24}$$

5. $\mathbf{F}^{\top}\mathbf{Q}\mathbf{Q}^{\top}\mathbf{F}$:

$$\partial = \mathbf{F} \mathbf{\Lambda} \mathbf{F}^{\mathsf{T}} \mathbf{Q} - \mathbf{Q}^{-\mathsf{T}}$$
$$\partial \langle \partial, \mathbf{M} \rangle = \mathbf{F} \mathbf{\Lambda} \mathbf{F}^{\mathsf{T}} \mathbf{M}^{\mathsf{T}} + \mathbf{Q}^{-\mathsf{T}} \mathbf{M}^{\mathsf{T}} \mathbf{Q}^{-\mathsf{T}}$$
(25)

In future notes, we will consider the full derivatives required for variational latent Gaussian-process inference for the Poisson and probit generalized linear models.